Paraphrases do not explain word analogies

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Abstract

Many types of distributional word embeddings (weakly) encode linguistic regularities as directions (the difference between *jump* and *jumped* will be in a similar direction to that of *walk* and *walked*, and so on). Several attempts have been made to explain this fact. We respond to Allen and Hospedales' recent (ICML, 2019) theoretical explanation, which claims that word2vec and GloVe will encode linguistic regularities whenever a specific relation of *paraphrase* holds between the four words involved in the regularity. We demonstrate that the explanation does not go through: the paraphrase relations needed under this explanation do not hold empirically.

1 Introduction

The study of linguistic regularities in distributional word embeddings—that the difference vector calculated between the vectors jump and jumped shows a similar direction to that of walk and walked, and so on—has been both stimulating and controversial. While a number of such regularities appear to hold, across a number of different kinds of embeddings, the standard 3COSADD analogy test used to measure the presence of these regularities has come under fire for confounding analogical regularities with unrelated properties of semantic embeddings. It is thus important to note that several papers have proposed theoretical explanations for why linguistic regularities should hold in distributional word embeddings. Particularly in light of the controversies over linguistic regularities, it is important to examine the soundness of these arguments.

Allen and Hospedales (2019) develop such an explanation by linking the semantic definition of an analogy to *paraphrases*. In the sense of Gittens et al. (2017), paraphrases are sets of words which are semantically and distributionally closely

equivalent to another word or set of words—for example, king may be paraphrased by $\{man, royal\}$. Allen and Hospedales argue that the standard analogy criterion, that king - man + woman = queen, is equivalent to a criterion whereby {king, woman} paraphrases $\{man, queen\}$. With this in mind, it becomes possible to rewrite the arithmetic analogy criterion in terms of vectors encoding the pointwise mutual information (PMI) between words and their contexts, and to decompose the error in the analogy equality into several components, including a paraphrase error term measuring the degree to which the critical paraphrase holds. Making use of an assumption that the word2vec embedding is a linear transformation of the PMI matrix, they argue that results in terms of PMI apply to word vectors. Thus, under their explanation, a major part of success on an analogy $a - a^* + b^* = b$ is due to a, b^* and a^* , b being close distributional paraphrases.

We first review the literature on the analogy test itself, underlining known pitfalls which any explanation of linguistic regularities must navigate. We then show empirically that the relation between the PMI matrix and word2vec embeddings is to some degree linear, which may be enough to satisfy the assumption of Allen and Hospedales (2019). We further examine the proposed decomposition into error terms. We demonstrate that, empirically, these error terms tend to be undefined due to data sparseness, undermining their explanatory force. Most importantly, examining a number of analogies which pass the standard test, we show that the critical paraphrase error term is, contrary to the proposed explanation, very large. ¹

¹Code is available at www.github.com/bootphon/paraphrases_do_not_explain_analogies.

2 Related work

Early works proposing explanations of the analogical properties of word embeddings include Mikolov et al. (2013b) and Pennington et al. (2014). A geometrical explanation is proposed by Arora et al. (2016), but this explanation relies on very strong preconditions, notably, that the word vectors be distributed uniformly in space. Ethayarajh et al. (2019) also propose an explanation, providing a link between the PMI and the norm of word embeddings. However, as pointed out by Allen and Hospedales (2019), this explanation, too, rests on strong assumptions. Notably, the words involved in the analogy are required to be coplanar, a property that seems unlikely in light of the lack of parallelism we discuss in the next section.

3 Issues with the test

Issues have arisen with the standard way of measuring linguistic analogies. Levy and Goldberg (2014a), Vylomova et al. (2016), Rogers et al. (2017), and Fournier et al. (2020) all demonstrate that the standard 3CosADD measure conflates several very different properties of embeddings, simultaneously measuring not only the directional regularities suggested by typical illustrations of vectors in a parallelogram, but also the similarity of individual matched pairs such as king, man, as well as the global arrangement of vectors in semantic fields, such as king, queen, prince, ... versus man, woman, child, ... in distinct regions of the space. These issues undermine the construct validity of the standard analogy test. This conflation of properties explains certain pathological behaviours of the test (Linzen, 2016; Rogers et al., 2017). In spite of these issues, Fournier et al. (2020) demonstrate, using alternative measures, that linguistic regularities are nevertheless coded by directional similarities. This parallelism is weak, with directions tending to be closer, in the absolute, to being orthogonal than to being parallel, but is present above chance level (unmatched word pairs).

Thus, before turning to Allen and Hospedales (2019), one of a number of theoretical attempts to explain performance on the 3COSADD objective, we underscore that such demonstrations run the risk of explaining properties of the test which may be of secondary interest, or, conversely, of placing undue emphasis on the role of directional regularities, which have been shown to play only a small role in success on 3COSADD.

4 Explaining analogies through paraphrases

For a word w_i and a word c_j which can appear in the context of w_i , the pairwise mutual information $PMI(w_i, c_j)$ is defined as $\log \frac{p(w_i, c_j)}{p(w_i)p(c_j)}$. As shown by Levy and Goldberg (2014b), skip-gram word2vec with negative sampling factorizes the PMI: PMI $\approx W^{\top} \cdot C$, with W and C the word and context embedding matrices of a word2vec model.

For two pairs of words (a, a^*) and (b, b^*) from the same semantic relation, the standard arithmetic analogy test criterion is that $a - a^* + b^* = b$. Writing $\mathcal{W} = \{a, b^*\}, \mathcal{W}_* = \{a^*, b\}$, and \mathbf{PMI}_x the PMI vector of x, Allen and Hospedales (2019) show that is possible to rewrite the arithmetic analogy formula with PMI vectors, and to decompose the error in the equality into five terms as follows:

$$\mathbf{PMI}_{b^*} = \mathbf{PMI}_b + \mathbf{PMI}_{a^*} - \mathbf{PMI}_a$$

$$+ \underbrace{\rho^{\mathcal{W}, \mathcal{W}_*}}_{\text{Paraphrase error}} + \underbrace{\sigma^{\mathcal{W}} - \sigma^{\mathcal{W}_*}}_{\text{Conditional dependence error}}$$

$$+ \underbrace{(\tau^{\mathcal{W}} - \tau^{\mathcal{W}_*})\mathbf{1}}_{\text{Mutual dependence error}}$$
(1)

The error terms are vectors of length $|\mathcal{V}|$ (vocabulary size), with each element j defined as:

$$\rho^{\mathcal{W}, \mathcal{W}_*} = \log \frac{p(c_j | \mathcal{W}_*)}{p(c_j | \mathcal{W})}$$

$$\sigma^{\mathcal{W}} = \log \frac{p(\mathcal{W} | c_j)}{\prod_{\mathcal{W}} p(w_i | c_j)}$$

$$\tau^{\mathcal{W}} = \log \frac{p(\mathcal{W})}{\prod_{\mathcal{W}} p(w_i)}$$
(2)

The authors claim that these terms can be embedded linearly into a word2vec embedding space by multiplying them by the Moore-Penrose pseudoinverse C^{\dagger} of the context matrix C. Then with \mathbf{w}_x the word2vec embedding of x, $C^{\dagger} \cdot \mathbf{PMI}_x \approx \mathbf{w}_x$. Thus we get the final decomposition:

$$\mathbf{w}_{b^*} = \mathbf{w}_b + \mathbf{w}_{a^*} - \mathbf{w}_a + C^{\dagger} \left(\rho^{\mathcal{W}, \mathcal{W}_*} + \sigma^{\mathcal{W}} - \sigma^{\mathcal{W}_*} - (\tau^{\mathcal{W}} - \tau^{\mathcal{W}_*}) \mathbf{1} \right)$$
(3)

The paraphrase error term is claimed to be small for successful analogies. Elaborating on the notation, W is taken to paraphrase W_* if, wherever

all $w \in \mathcal{W}$ appear together, we observe the same distribution of surrounding words as for \mathcal{W}_* . The paraphrase error assesses the similarity of the distributions of words in the context of \mathcal{W} (all words in \mathcal{W} appearing together) versus \mathcal{W}_* .

5 Linearity of the link between PMI and word2vec

Though it is true that there is a relation between the word2vec matrices $W^{\top} \cdot C$ and the PMI matrix, in practice the link is more complicated than simple linear matrix factorization, due in part to the training tricks described in Mikolov et al. (2013a). The result of Allen and Hospedales (2019) requires that the embedding from PMI vectors to word2vec embeddings be "linear enough" for C^{\dagger} · PMI to approximate W.

To assess this, we use the *text8* corpus 2 both to train word2vec embeddings 3 and to estimate a PMI matrix. We replace infinite values in the PMI matrix by 0. In Figure 1a, we show the distribution of the Pearson correlation coefficient (assessing the presence of a linear relation) between the word2vec embedding and the corresponding row of C^{\dagger} · PMI for the top ten thousand words in the corpus. As can be seen from the figure, the correlation tends to be between 0.5 and 0.8. For instance in Figure 1b, the word2vec embedding for *king* is plotted against the row of C^{\dagger} · PMI corresponding to *king*.

While the relation is not perfectly linear—many words have a correlation of around 0.55, far lower than that of *king*—the empirical relations shown here leave open the possibility that it may indeed be "sufficiently linear" to be taken for granted. However, while linearity is necessary for the result of Allen and Hospedales (2019) to go through, it is not sufficient. In the next section, we assess the critical question of whether the paraphrase error is small enough to serve as an explanation for the success of linguistic analogies.

6 Empirical analysis of the error terms

We now seek to examine the proposed explanation by calculating the proposed error terms empirically. However, in practice, many of the terms

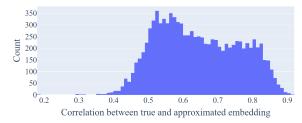


Figure 1a: Histogram of the Pearson correlations between true and approximated word2vec embeddings for the top ten thousand words in the *text8* corpus. The mean value is 0.643 and the variance is 0.014.

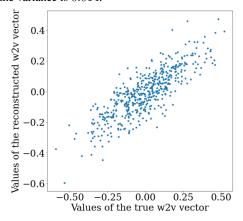


Figure 1b: Plot of the values of the word2vec embedding for king, versus coefficients for the row of $C^{\dagger} \cdot PMI$ corresponding to king, for word2vec matrices trained on the same corpus (text8). The Pearson correlation is one of the best possible at 0.825.

are undefined, since they rely on cooccurrences unattested in practical corpora. The most extreme situation occurs when the two words of a paraphrase $\mathcal{W} = \{w_1, w_2\}$ are never present in the same context window in the corpus. We found that only 16% of the paraphrase sets associated with the BATS analogy set (Gladkova et al., 2016)—for example, king, woman—were present together in the text8 corpus in a context window of length five. We refer to such paraphrase sets as "well-defined" with respect to the corpus. The problem of zero co-occurrence counts was anticipated by Allen and Hospedales (2019), who propose to restrict their analysis to the case where the context window is sufficiently large that all relevant terms are well defined. We stress that our trained word2vec vectors are also trained with a context window of five, and yield expected levels of performance on the BATS analogy test, despite having access to little training data on which to model co-occurrences such as king, woman, queen, man, and so on.

At a minimum, if the proposed explanation holds, the cases for which the error terms are empirically well-defined should show signs of the paraphrase

²A text dataset composed of 100 million characters from Wikipedia: (Mahoney, 2006).

³Skip-gram architecture with negative sampling (1 word), negative sampling exponent equal to 1, no undersampling of common words, and a high dimension size of 500. These parameters allow us to be as close as possible to a direct factorization of the PMI matrix.

Category	I01	102	I05	I06	I07	I08	I09	I10	D02	D03	D05	D08	D10	E01	E02
Paraphrase error norm	177	153	111	127	126	124	138	97	102	122	130	110	107	124	176
Dependence errors sum norm	1006	938	867	903	957	883	952	908	856	893	514	585	699	749	848
All errors sum norm	1032	957	878	917	970	897	966	916	864	905	539	602	710	765	875
Category	E03	E04	E05	E08	E09	E10	L02	L03	L04	L05	L06	L07	L08	L09	L10
Paraphrase error norm	162	176	155	229	179	190	197	189	209	206	133	169	185	175	432
Dependence errors sum norm	866	797	519	739	910	833	642	982	907	1103	921	995	1044	1017	1302
All errors sum norm	889	822	553	778	933	865	683	1007	939	1131	937	1016	1066	1040	1416

Table 1: L2 norms of the error terms in 1, following our implementation.

Category	I01	I02	105	I06	I07	108	109	I10	D02	D03	D05	D08	D10	E01	E02
Average rank	7762K	7589K	7759K	8744K	8160K	6454K	7028K	11889K	31952K	19558K	7857K	1506K	2556K	4394K	9507K
Median rank	1630K	2195K	3055K	3239K	2530K	4090K	3004K	4535K	6754K	3564K	3260K	1506K	2556K	2117K	1622K
Category	E03	E04	E05	E08	E09	E10	L02	L03	L04	L05	L06	L07	L08	L09	L10
Category Average rank	E03 1305K	E04 5611K	E05 9192K	E08 727K	E09 8421K	E10 11946K	L02 52183K	L03 1857K	L04 12687K	L05 6747K	L06 2475K	L07 7727K	L08 4502K	L09 4679K	L10 16871K

Table 2: For an analogy equivalent to two paraphrases W and W_* , the rank of W_* in the list of the closest paraphrases to W with respect to the L2 norm of the paraphrase error vector. 7762K means a rank of 7762000, rounded to the nearest thousand.

error being relatively small. We now detail how we implemented the error terms in cases for which they were well-defined. We count co-occurrences $N(w_i, w_j, w_k)$ in text8 for all triplets of words w_i, w_i, w_k , with w_k at the center of the context window, and $W = \{w_i, w_i\}$ any paraphrase, both occurring anywhere within a context window of width five. We restrict analysis to the ten thousand most frequent word types w_i and w_j , yielding 10^8 possible paraphrases.⁴ We use the relative frequencies as estimators of $p(w_k|\{w_i, w_i\})$ and $p(\{w_i, w_i\}|w_k)$, and marginalize to obtain $p(w_i|w_k)$, $p(\{w_i, w_i\})$ and $p(w_k)$. The error terms follow. Since this can still lead to ill-defined elements, we replace $log(+\infty)$ and log(0) by $+/-log(\epsilon)$, with $\epsilon =$ 10^{-15} (within reason, the value of ϵ is immaterial). We also replace log(0/0) with 0.

Table 1 shows the mean and median values of the L2 norms of the paraphrase error vectors across several categories of the BATS dataset. We compare them with the sum of the four dependence error terms (the dependence error reflects statistical dependencies within \mathcal{W} and \mathcal{W}_* irrelevant to the analogy), as well as the sum of all five error terms (equal to the difference between the PMI of \mathcal{W} and \mathcal{W}_*). The paraphrase error is indeed smaller than the other error terms. However, as we now show, the paraphrase error is not small *enough* to contribute substantially to the success of analogies. ⁵

Take the norm of the paraphrase error vector ρ as a measure of the divergence in the PMI between two paraphrases. For an analogy with associated paraphrases W and W_* , we assess how many paraphrases are closer to ${\mathcal W}$ than to ${\mathcal W}_*$ by calculating the rank of the norm of ρ^{W,W_*} among all $\rho^{W,X}$, where X spans over all pairs of words constructible from the top ten thousand most frequent words in the corpus. To do so, we define a Paraphrase Conditional Information matrix (PCI). For $W_{ij} = \{w_i, w_j\}$ and w_k , we define $PCI(l_{ij}, k)$, the value at column l_{ij} and row k to be $log(p(W_{ij}|w_k))$, where with l_{ij} is a unique index associate with tuple (i, j). We compute only the positive PCI, to obtain a sparse matrix. The difference between two PCI columns is a paraphrase error vector, and their Euclidean distance is the norm of the paraphrase error.

We now compute, for each analogy, the distance between the PCI column of \mathcal{W} and every other column (paraphrase) of the PCI matrix. We calculate the rank of the true analogy pair \mathcal{W}_* . Given that the analogy test generally succeeds in picking out b as being the most similar to $a-a^*+b^*$ out of the entire vocabulary (modulo Linzen 2016), we would expect that, for successful analogies, the paraphrase error for the true analogy would be among the highest, if small paraphrase error were the explanation for success. Table 2 displays the mean of this rank within each BATS category. The rank is extremely low (in the millions), making the paraphrase error in true analogies far too high to be the explanation for their success.

 $^{^4}w_k$ is allowed to vary over all of the types included in the training for word2vec, of which there are 71290. Thus, for each paraphrase, the error vectors have 71290 elements, one for each vocabulary word.

⁵We note also that the error values seem relatively consistent between categories, while success on the analogy test varies differ greatly between categories.

 $^{^6}$ Limiting the search to the paraphrases composed by at least one of the words of W_* still results in a very low rank

7 Conclusion

Recent work has shown that, in spite of the standard analogy test's confound with simple vector similarity, distributional word vectors genuinely do encode linguistic regularities as directional regularities above and beyond vector similarity (Fournier et al., 2020), further research is warranted into the mechanisms by which distributional word embeddings come to show these regularities. However, the analysis of analogies as paraphrases does not hold up as an explanation of performance on the analogy test—nor would an explanation of performance on the 3COSADD analogy test be a satisfying result, since the test is not a useful measure to begin with.

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